Random Unitary Matrices and Friends

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- Metric Structure:
 - ► U(n) sits inside C^{n²} and inherits a geodesic metric d_g(·, ·)
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Randomness:

There is a unique translation-invariant probability measure called Haar measure on $\mathbb{U}(n)$: if *U* is a Haar-distributed random unitary matrix, so are *AU* and *UA*, for *A* a fixed unitary matrix.

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1. Pick the first column U_1 uniformly from $\mathbb{S}^1_{\mathbb{C}} \subseteq \mathbb{C}^n$.

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- Fill an n × n array with i.i.d. standard complex Gaussian random variables.
 - Stick the result into the *QR* algorithm; the resulting *Q* is Haar-distributed on U(*n*).

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- The set of all *n* × *n* unitary matrices is denoted (*n*); this set is a subgroup and a submanifold of U (*n*).
- O (*n*) has two connected components: SO (*n*) (det(U) = 1) and SO[−] (*n*) (det(U) = −1).
- ► There is a unique translation-invariant (Haar) probability measure on each of O (n), SO (n) and SO⁻ (n).

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► A symplectic matrix is an 2n × 2n matrix with entries in C, such that

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• The group of $2n \times 2n$ symplectic matrices is denoted $\mathbb{S}_{\mathbb{P}}(2n)$.

Concentration of measure

Theorem (G/M;B/E;L;M/M)

Let G be one of SO(n), $SO^{-}(n)$, SU(n), U(n), $S_{\mathbb{P}}(2n)$, and let $F: G \to \mathbb{R}$ be L-Lipschitz (w.r.t. the geodesic metric or the HS-metric). Let U be distributed according to Haar measure on G. Then there are universal constants C, c such that

$\mathbb{P}\left[\left|F(U)-\mathbb{E}F(U)\right|>Lt\right]\leq Ce^{-cnt^2},$

for every t > 0.

Note: permuting the rows or columns of a random orthogonal matrix *U* corresponds to left- or right-multiplication by a permutation matrix (which is itself orthogonal).

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Classical fact: A coordinate of a random point on the sphere in \mathbb{R}^n is approximately Gaussian, for large *n*.

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if U is large.

A more modern fact (Diaconis–Freedman): If X is a randomly distributed point on the sphere of radius \sqrt{n} in \mathbb{R}^n , and Z is a standard Gaussian random vector in \mathbb{R}^n , then

$$d_{TV}\Big((X_1,\ldots,X_k),(Z_1,\ldots,Z_k)\Big) \leq \frac{2(k+3)}{n-k-3}$$

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Diaconis'question: How many entries of *U* can be simultaneously approximated by independent Gaussians?

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Theorem (Jiang)

Let $\{U_n\}$ be a sequence of random orthogonal matrices with $U_n \in \mathbb{O}(n)$ for each *n*, and suppose that $p_n, q_n = o(\sqrt{n})$.

Let $\mathcal{L}(\sqrt{n}U(p_n, q_n))$ denote the joint distribution of the p_nq_n entries of the top-left $p_n \times q_n$ block of $\sqrt{n}U_n$, and let $Z(p_n, q_n)$ denote a collection of p_nq_n i.i.d. standard normal random variables. Then

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$\lim_{n\to\infty} d_{TV}(\mathcal{L}(\sqrt{n}U(p_n,q_n)),Z(p_n,q_n))=0.$

That is, a $p_n \times q_n$ principle submatrix can be approximated in total variation by a Gaussian random matrix, as long as $p_n, q_n \ll \sqrt{n}$.
Jiang's answer(s)

Theorem (Jiang) For each *n*, let $Y_n = [y_{ij}]_{i,j=1}^n$ be an $n \times n$ matrix of independent standard Gaussian random variables and let $\Gamma_n = [\gamma_{ij}]_{i,j=1}^n$ be the matrix obtained from Y_n by performing the Gram-Schmidt process; i.e., Γ_n is a random orthogonal matrix. Let

$$\epsilon_n(m) = \max_{1 \le i \le n, 1 \le j \le m} \big| \sqrt{n} \gamma_{ij} - \mathbf{y}_{ij} \big|.$$

Then

$$\epsilon_n(m_n) \xrightarrow[n \to \infty]{\mathbb{P}} 0$$

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Then

$$\epsilon_n(m_n) \xrightarrow[n \to \infty]{\mathbb{P}} 0$$

if and only if $m_n = o\left(\frac{n}{\log(n)}\right)$.

That is, in an "in probability" sense, $\frac{n^2}{\log(n)}$ entries of *U* can be simultaneously approximated by independent Gaussians.

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Choosing a principle submatrix of an $n \times n$ orthogonal matrix U corresponds to a particular type of orthogonal projection from a large matrix space to a smaller one.

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(Note that the result is no longer orthogonal.)

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In general, a rank k orthogonal projection of $\mathbb{O}(n)$ looks like

$$U \mapsto (\operatorname{Tr}(A_1 U), \ldots, \operatorname{Tr}(A_k U)),$$

where A_1, \ldots, A_k are orthonormal matrices in $\mathbb{O}(n)$; i.e.,

$$\operatorname{Tr}(\boldsymbol{A}_{i}\boldsymbol{A}_{j}^{T})=\delta_{ij}.$$

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Theorem (Chatterjee–M.)

Let A_1, \ldots, A_k be orthonormal (w.r.t. the Hilbert-Schmidt inner product) in $\mathbb{O}(n)$, and let $U \in \mathbb{O}(n)$ be a random orthogonal matrix. Consider the random vector

 $X := (\mathrm{Tr}(A_1 U), \ldots, \mathrm{Tr}(A_k U)),$

and let $Z := (Z_1, ..., Z_k)$ be a standard Gaussian random vector in \mathbb{R}^k . Then for all $n \ge 2$,

$$d_W(X,Z) \leq \frac{\sqrt{2}k}{n-1}.$$

Here, $d_W(\cdot, \cdot)$ denotes the L₁-Wasserstein distance.

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Let *U* be a Haar-distributed matrix in $\mathbb{U}(N)$.

Then *U* has (random) eigenvalues $\{e^{i\theta_j}\}_{j=1}^N$.

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Note: The distribution of the set of eigenvalues is rotation-invariant.

To understand the behavior of the ensemble of random eigenvalues, we consider the empirical spectral measure of *U*:

$$\mu_N := \frac{1}{N} \sum_{j=1}^N \delta_{e^{i\theta_j}}.$$

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100 i.i.d. uniform random points



E. Rains

The eigenvalues of a 100×100 random unitary matrix

Diaconis/Shahshahani

Theorem (D–S)

Let $U_n \in \mathbb{U}(n)$ be a random unitary matrix, and let μ_{U_n} denote the empirical spectral measure of U_n . Let ν denote the uniform probability measure on \mathbb{S}^1 . Then

$$u_{U_n} \xrightarrow{n \to \infty} \nu,$$

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- ► The theorem follows from explicit formulae for the mixed moments of the random vector (Tr(U_n),...,Tr(U_n^k)) for fixed k, which have been useful in many other contexts.
- ► They showed in particular that (Tr(U_n),...,Tr(U^k_n)) is asymptotically distributed as a standard complex Gaussian random vector.

The number of eigenvalues in an arc

Theorem (Wieand)

Let $I_j := (e^{i\alpha_j}, e^{i\beta_j})$ be intervals on \mathbb{S}^1 and for $U_n \in \mathbb{U}(n)$ a random unitary matrix, let

$$Y_{n,k} := \frac{\mu_{U_n}(I_k) - \mathbb{E}\mu_{U_n}(I_k)}{\frac{1}{\pi}\sqrt{\log(n)}}.$$

Then as n tends to infinity, the random vector $(Y_{n,1}, \ldots, Y_{n,k})$ converges in distribution to a jointly Gaussian random vector (Z_1, \ldots, Z_k) with covariance

$$\operatorname{Cov}(Z_j, Z_k) = \begin{cases} 0, & \alpha_j, \alpha_k, \beta_j, \beta_k \text{ all distict;} \\ \frac{1}{2} & \alpha_j = \alpha_k \text{ or } \beta_j = \beta_k \text{ (but not both);} \\ -\frac{1}{2} & \alpha_j = \beta_k \text{ or } \beta_j = \alpha_k \text{ (but not both);} \\ 1 & \alpha_j = \alpha_k \text{ and } \beta_j = \beta_k; \\ -1 & \alpha_j = \beta_k \text{ and } \beta_j = \alpha_k. \end{cases}$$

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About that weird covariance structure...

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Another Gaussian process that has it:

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Another Gaussian process that has it: Again suppose that $I_j := (e^{i\alpha_j}, e^{i\beta_j})$ are intervals on \mathbb{S}^1 , and suppose that $\{G_{\theta}\}_{\theta \in [0,2\pi)}$ are i.i.d. standard Gaussians. Define

$$X_{n,k} = G_{\beta_k} - G_{\alpha_k};$$

then

$$\operatorname{Cov}(X_j, X_k) = \begin{cases} 0, & \alpha_j, \alpha_k, \beta_j, \beta_k \text{ all distict;} \\ \frac{1}{2} & \alpha_j = \alpha_k \text{ or } \beta_j = \beta_k \text{ (but not both);} \\ -\frac{1}{2} & \alpha_j = \beta_k \text{ or } \beta_j = \alpha_k \text{ (but not both);} \\ 1 & \alpha_j = \alpha_k \text{ and } \beta_j = \beta_k; \\ -1 & \alpha_j = \beta_k \text{ and } \beta_j = \alpha_k. \end{cases}$$

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Where's the white noise in *U*?

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Theorem (Hughes–Keating–O'Connel)

Let $Z(\theta)$ be the characteristic polynomial of U and fix $\theta_1 \dots, \theta_k$. Then

$$\frac{1}{\sqrt{\frac{1}{2}\log(n)}} \left(\log(Z(\theta_1)), \dots, \log(Z(\theta_k))\right)$$

converges in distribution to a standard Gaussian random vector in \mathbb{C}^k , as $n \to \infty$.

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HKO in particular showed that Wieand's result follows from theirs by the argument principle.

Powers of U



The eigenvalues of U^m for m = 1, 5, 20, 45, 80, for U a realization of a random 80×80 unitary matrix.

Rains' Theorems

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Theorem (Rains 1997)

Let $U \in \mathbb{U}(n)$ be a random unitary matrix, and let $m \ge n$. Then the eigenvalues of U^m are distributed exactly as n i.i.d. uniform points on \mathbb{S}^1 .

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Theorem (Rains 2003) Let $m \le N$ be fixed. Then

$$\left[\mathbb{U}(N)\right]^{m} \stackrel{e.v.d.}{=} \bigoplus_{0 \leq j < m} \mathbb{U}\left(\left\lceil \frac{N-j}{m} \right\rceil\right),$$

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where $\stackrel{e.v.d.}{=}$ denotes equality of eigenvalue distributions.



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Theorem (E.M./M. Meckes)

Let ν denote the uniform probability measure on the circle and $W_p(\mu,\nu) := \inf \left\{ \left(\int |x-y|^p d\pi(x,y) \right)^{\frac{1}{p}} \middle| \begin{array}{l} \pi(A \times \mathbb{C}) = \mu(A) \\ \pi(\mathbb{C} \times A) = \nu(A) \end{array} \right\}.$

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$$\mathbb{E}\left[W_{\rho}(\mu_{m,N},\nu)\right] \leq \frac{C\rho\sqrt{m\left[\log\left(\frac{N}{m}\right)+1\right]}}{N}$$

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Almost sure convergence

Corollary

For each N, let U_N be distributed according to uniform measure on $\mathbb{U}(N)$ and let $m_N \in \{1, ..., N\}$. There is a C such that, with probability 1,

$$W_{p}(\mu_{m_{N},N},
u) \leq rac{Cp\sqrt{m_{N}\log(N)}}{N^{rac{1}{2}+rac{1}{\max(2,p)}}}$$

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eventually.

A miraculous representation of the eigenvalue counting function

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A miraculous representation of the eigenvalue counting function

Fact: The set $\{e^{i\theta_j}\}_{j=1}^N$ of eigenvalues of U (uniform in $\mathbb{U}(N)$) is a determinantal point process.

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A miraculous representation of the eigenvalue counting function

Fact: The set $\{e^{i\theta_j}\}_{j=1}^N$ of eigenvalues of U (uniform in $\mathbb{U}(N)$) is a determinantal point process.

Theorem (Hough/Krishnapur/Peres/Virág 2006)

Let \mathcal{X} be a determinantal point process in Λ satisfying some niceness conditions. For $D \subseteq \Lambda$, let \mathcal{N}_D be the number of points of \mathcal{X} in D. Then

$$\mathcal{N}_D \stackrel{d}{=} \sum_k \xi_k,$$

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where $\{\xi_k\}$ are independent Bernoulli random variables with means given explicitly in terms of the kernel of \mathcal{X} .
A miraculous representation of the eigenvalue counting function

That is, if \mathcal{N}_{θ} is the number of eigenangles of *U* between 0 and θ , then

$$\mathcal{N}_{\theta} \stackrel{d}{=} \sum_{j=1}^{N} \xi_{j}$$

for a collection $\{\xi_j\}_{j=1}^N$ of independent Bernoulli random variables.

A miraculous representation of the eigenvalue counting function

Recall Rains' second theorem:

$$\left[\mathbb{U}(N)\right]^{m} \stackrel{e.v.d.}{=} \bigoplus_{0 \leq j < m} \mathbb{U}\left(\left\lceil \frac{N-j}{m} \right\rceil\right),$$

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$$\left[\mathbb{U}(N)\right]^{m} \stackrel{e.v.d.}{=} \bigoplus_{0 \leq j < m} \mathbb{U}\left(\left\lceil \frac{N-j}{m} \right\rceil\right),$$

So: if $\mathcal{N}_{m,N}(\theta)$ denotes the number of eigenangles of U^m in $[0, \theta)$, then

$$\mathcal{N}_{m,N}(\theta) \stackrel{d}{=} \sum_{j=1}^{N} \xi_j,$$

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for $\{\xi_j\}_{j=1}^N$ independent Bernoulli random variables.

► From Bernstein's inequality and the representation of $\mathcal{N}_{m,N}(\theta)$ as $\sum_{j=1}^{N} \xi_j$,

$$\mathbb{P}\left[\left|\mathcal{N}_{m,N}(\theta)-\mathbb{E}\mathcal{N}_{m,N}(\theta)\right|>t\right]\leq 2\exp\left[-\min\left\{\frac{t^2}{4\sigma^2},\frac{t}{2}\right\}\right],$$

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where $\sigma^2 = \operatorname{Var} \mathcal{N}_{m,N}(\theta)$.

From Bernstein's inequality and the representation of $\mathcal{N}_{m,N}(\theta)$ as $\sum_{j=1}^{N} \xi_j$,

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where $\sigma^2 = \operatorname{Var} \mathcal{N}_{m,N}(\theta)$.

• $\mathbb{E}\mathcal{N}_{m,N}(\theta) = \frac{N\theta}{2\pi}$ (by rotation invariance).

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where $\sigma^2 = \operatorname{Var} \mathcal{N}_{m,N}(\theta)$.

- $\mathbb{E}\mathcal{N}_{m,N}(\theta) = \frac{N\theta}{2\pi}$ (by rotation invariance).
- Var [N_{1,N}(θ)] ≤ log(N) + 1 (e.g., via explicit computation with the kernel of the determinantal point process), and so

$$\operatorname{Var}\left(\mathcal{N}_{m,N}(\theta)\right) = \sum_{0 \le j < m} \operatorname{Var}\left(\mathcal{N}_{1,\left\lceil \frac{N-j}{m} \right\rceil}(\theta)\right) \le m\left(\log\left(\frac{N}{m}\right) + 1\right).$$

$$\mathbb{P}\left[\left|\theta_j - \frac{2\pi j}{N}\right| > \frac{4\pi t}{N}\right] \le 4 \exp\left[-\min\left\{\frac{t^2}{m\left(\log\left(\frac{N}{m}\right) + 1\right)}, t\right\}\right],$$

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for each $j \in \{1, ..., N\}$:

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for each $j \in \{1, ..., N\}$:

$$\mathbb{P}\left[\theta_j > \frac{2\pi j}{N} + \frac{4\pi}{N}u\right] = \mathbb{P}\left[\mathcal{N}_{\frac{2\pi(j+2u)}{N}}^{(m)} < j\right]$$

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for each $j \in \{1, ..., N\}$:

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$$= \mathbb{P}\left[j + 2u - \mathcal{N}_{\frac{2\pi(j+2u)}{N}}^{(m)} > 2u\right]$$

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$$\mathbb{P}\left[\left|\theta_j - \frac{2\pi j}{N}\right| > \frac{4\pi t}{N}\right] \le 4 \exp\left[-\min\left\{\frac{t^2}{m\left(\log\left(\frac{N}{m}\right) + 1\right)}, t\right\}\right],$$

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Bounding $\mathbb{E} W_{p}(\mu_{m,N},\nu)$

If
$$\nu_N := \frac{1}{N} \sum_{j=1}^N \delta_{\exp\left(j\frac{2\pi j}{N}\right)}$$
, then $W_p(\nu_N, \nu) \le \frac{\pi}{N}$ and

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Bounding $\mathbb{E} W_{p}(\mu_{m,N},\nu)$

If
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, then $W_p(\nu_N, \nu) \le \frac{\pi}{N}$ and

$$\mathbb{E} \textit{W}^{\textit{p}}_{\textit{p}}(\mu_{\textit{m},\textit{N}},\nu_{\textit{N}}) \leq \frac{1}{\textit{N}} \sum_{j=1}^{\textit{N}} \mathbb{E} \left| \theta_j - \frac{2\pi j}{\textit{N}} \right|^{\textit{p}}$$

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 $\mathbb{E}W_p^p(\mu_{m,N}, \nu_N) \le \frac{1}{N} \sum_{j=1}^N \mathbb{E} \left| \theta_j - \frac{2\pi j}{N} \right|^p$

$$\leq 8\Gamma(
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using the concentration result and Fubini's theorem.

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The Idea: Consider the function $F_p(U) = W_p(\mu_{U^m}, \nu)$, where μ_{U^m} is the empirical spectral measure of U^m .

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► By Rains' theorem, it is distributionally the same as $F_p(U_1, ..., U_m) = \left(\frac{1}{m} \sum_{j=1}^m \mu_{U_j}, \nu\right).$

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- ► F_p(U₁,..., U_m) is Lipschitz (w.r.t. the L₂ sum of the Euclidean metrics) with Lipschitz constant N^{- ¹max(p,2)}.
- ► If we had a general concentration phenomenon on $\bigoplus_{0 \le j < m} \mathbb{U}\left(\left\lceil \frac{N-j}{m} \right\rceil\right)$, concentration of $W_p(\mu_{U^m}, \nu)$ would follow.

Concentration on $\mathbb{U}(N_1) \oplus \cdots \oplus \mathbb{U}(N_k)$

Theorem (E. M./M. Meckes)

Given $N_1, \ldots, N_k \in \mathbb{N}$, denote by $M = \mathbb{U}(N_1) \times \cdots \mathbb{U}(N_k)$ equipped with the L_2 -sum of Hilbert–Schmidt metrics.

Suppose that $F : M \to \mathbb{R}$ is L-Lipschitz, and that $U_j \in \mathbb{U}(N_j)$ are independent, uniform random unitary matrices, for $1 \le j \le k$. Then for each t > 0,

$$\mathbb{P}\Big[F(U_1,\ldots,U_k)\geq \mathbb{E}F(U_1,\ldots,U_k)+t\Big]\leq e^{-Nt^2/12L^2},$$

where $N = \min\{N_1, ..., N_k\}$.